Use of the CRIB (clinical risk index for babies) score in prediction of neonatal mortality and morbidity


Abstract
A prospective study of the outcome of care of a regional cohort of very low birthweight (<1500 g) and very preterm (<32 weeks) infants was carried out. Its aims were to assess the ability of the CRIB (clinical risk index for babies) score, rather than gestational age or birthweight, to predict mortality before hospital discharge, neurological morbidity, and length of stay, and to access CRIB score as an indicator of neonatal intensive care performance. 676 live births fulfilled the criteria and complete data were available for 643 (95%). Compared with gestation and birthweight, CRIB was better for the prediction of mortality, was as good for the prediction of morbidity, and was not as good for the prediction of length of stay. CRIB adjusted mortality did not demonstrate better performance in units providing the highest level of care. Either the CRIB score was not sensitive to performance or the level 3 hospitals in this study were performing badly.

On the basis of this analysis purchasers and providers of neonatal intensive care cannot yet rely on the CRIB score as a performance indicator.

(Keywords: CRIB score, very low birthweight infant, outcome.)

Health care professionals and purchasers of health care need to assess the requirements for neonatal intensive care, and to monitor the performance of units providing neonatal services. The only routinely available data in the United Kingdom are neonatal and birthweight specific mortality rates which take no account of case-mix and are not specific to neonatal care. Several scores have been published which use clinical details recorded at the time of admission to the neonatal unit and physiological variables recorded within the first 24 hours of life. The CRIB (clinical risk index for babies) score has been reported to be a better predictor of mortality than birthweight, and to correlate with morbidity at discharge from hospital, as measured by ultrasound evidence of major cerebral abnormality. The CRIB score has also been reported to take account of case-mix and to adjust crude mortality rates to allow comparison of the performance of neonatal intensive care units. After such an adjustment, teaching hospitals performed better than non-teaching hospitals. The study was criticised for failing to define teaching hospitals; doubts were expressed about the ability of CRIB to rank hospital performance and bias in case selection was suspected.

This study examines the outcome of care of a regional cohort of very low birthweight (<1500 g) and very preterm (<32 weeks) infants. Its aims were to assess: (1) the ability of the CRIB score, compared with gestational age and birthweight, to predict mortality before hospital discharge, neonatal morbidity, and length of stay; and (2) the use of the CRIB score as an indicator of neonatal intensive care performance.

Methods
The South East Thames Low Birthweight Study collected data on all infants with a birthweight between 500 and 2499 g and born to mothers resident in the region between 1 September 1992 and 31 August 1993. Babies weighing less than 500 g were included if gestational age was 22 weeks or more. All data were collected prospectively from each of the hospitals serving regional residents by two researchers (R de C-W, MS). The total study population was 3456 infants, comprising 95% of births notified to the Regional Child Health Computer database. The present analysis was limited to liveborn very low birthweight (<1500 g) and very preterm (<32 weeks) infants admitted to the special care baby unit (SCBU) and for whom the CRIB score would be relevant.

Six factors necessary for calculation of the CRIB score were collected — that is, birthweight, gestational age, the presence of congenital malformation(s) and maximum base excess, minimum and maximum appropriate inspired oxygen concentration in the first 12 hours. Mortality was defined as a death occurring before discharge from hospital. Neonatal neurological morbidity was defined as evidence of intraventricular haemorrhage (Papile grades 3 or 4) or cranial ultrasound scan. The CRIB score was divided into four subgroups: 0–5, 6–10, 11–15 and more than 153. Gestation was rounded to 23 weeks or less, 32 weeks or more, and to whole weeks in between. Birthweight was subdivided into six groups: less than 500 g, 500–749, 750–999, 1000–1249, 1250–1499, and 1500+. The relations of CRIB score, gestation, and
birthweight to hospital mortality for all babies, and to morbidity for survivors were analysed using univariate analysis. Spearman's rank correlation coefficient was used to explore their association with length of stay.

Receiver operating characteristic (ROC) curves compared the performance of different tests, by plotting sensitivity or detection rate against 1-specificity (false positive rate). The best test is one which achieves the highest sensitivity for the lowest false positive rate. ROC curves were composed for CRIB score, birthweight, and gestation to assess the ability of each to predict hospital mortality and, among survivors, neurological morbidity.

Babies were allocated to the hospital which provided most of their care in the first 72 hours of life. Hospitals 1–18 were all within South East Thames region. Hospital number 19 comprised data from several hospitals out of the region in which babies of resident mothers were delivered or to which they were transferred. In the analysis of their performance, hospitals 1–18 were categorised in three ways (table 1): level of care; workload; and designation.

Paneth et al described three levels of neonatal intensive care with level 1 being the highest level of care. Field et al related workload to the number of hours of ventilation per year, which is equivalent to category 1 maximal intensive care, in addition to specialist neonatal consultant input and 24 hour middle grade medical cover. A high workload was defined as more than 500 days of category 1 care per year. South East Thames Regional Health Authority has designated five neonatal units to be tertiary referral neonatal centres.

Crude hospital mortality rates were adjusted for CRIB using multiple logistic regression analysis to compare the performance of individual neonatal units, as described by the International Neonatal Network, and between the three different categories of units. The rankings of the individual units were compared with indirect standardised mortality ratios. The standardised mortality ratio (SMR) is described by the equation:

$$\text{SMR} = \frac{n \times \text{expected}}{\text{observed}}$$

where n is the number of babies, E is the expected mortality rate, and o is the observed number of deaths among all regional births in each of the four subgroups (0–5, 6–10, 11–15, 16+) of the CRIB score.

The data from this study were used to recreate the CRIB score by repeating the original methodology. This provided new values for the components of the six factors in the score and improvements to the score were then tested.

## Results

There were 676 babies admitted, of less than 32 weeks of gestation or weighing less than 1500 g. Analysis was restricted to the 643 with complete CRIB data. Table 2 shows the numbers of babies treated, numbers of deaths, and mean CRIB scores in each hospital. There were significantly more deaths among the cases with an incomplete CRIB score (13/33 = 39.4% vs 101/643 = 15.7%; χ² = 12.56; P=0.0001). However, six of the deaths were in babies admitted for terminal care. The other seven lived for 12 hours or more, had respiratory support, and would have been expected to have had a blood gas analysis performed, but some data were missing. After excluding the six babies admitted for terminal care there was no longer a significant difference in mortality (7/27 = 25.9% vs 101/643 = 15.7%; χ² = 1.338; P=0.09).

CRIB score, gestation, and birthweight were all significant univariate predictors of hospital mortality (P<0.0001). For example, mortality in hospital rose from 3% (14/464) with a CRIB score of 0–5 to 94% (15/16) with a score >15. Mortality in hospital rose from 7% (9/129) in babies born at 31 weeks' gestation to 61% (11/18) at 24 weeks. Mortality in hospital rose from 6% (9/152) in those whose birthweight was >1249 g to 73% (70/96) in those weighing <750 g. The ROC curve showed that CRIB score predicted mortality with greater sensitivity, at all levels of specificity, than did gestation or birthweight (fig 1).

**Table 1**

<table>
<thead>
<tr>
<th>Hospital</th>
<th>Admissions to SCBU*</th>
<th>Workload*</th>
<th>Category 1 days/year</th>
<th>Category 2 days/year</th>
<th>Regionally designated NICU</th>
<th>Penalty level of care</th>
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</table>

*Data for admissions and workload for the year 1990 taken from the annual report of the Regional Perinatal Monitoring Group. †Admission data for hospitals 3 and 4, 12 and 13, and 11 and 14 combined. ‡High workload; more than 500 hours of category 1 intensive care a year.

## Results

There were 676 babies admitted, of less than 32 weeks of gestation or weighing less than 1500 g. Analysis was restricted to the 643 with complete CRIB data. Table 2 shows the numbers of babies treated, numbers of deaths, and mean CRIB scores in each hospital. There were significantly more deaths among the cases with an incomplete CRIB score (13/33 = 39.4% vs 101/643 = 15.7%; χ² = 12.56; P=0.0001). However, six of the deaths were in babies admitted for terminal care. The other seven lived for 12 hours or more, had respiratory support, and would have been expected to have had a blood gas analysis performed, but some data were missing. After excluding the six babies admitted for terminal care there was no longer a significant difference in mortality (7/27 = 25.9% vs 101/643 = 15.7%; χ² = 1.338; P=0.09).

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**Table 2**

<table>
<thead>
<tr>
<th>Hospital</th>
<th>Total</th>
<th>Deaths</th>
<th>Mean CRIB score</th>
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<tbody>
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<td>1</td>
<td>74</td>
<td>13</td>
<td>3.5</td>
</tr>
<tr>
<td>2</td>
<td>103</td>
<td>21</td>
<td>4.4</td>
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<td>3</td>
<td>36</td>
<td>8</td>
<td>4.0</td>
</tr>
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<td>4</td>
<td>45</td>
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<td>21</td>
<td>1</td>
<td>1.6</td>
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<tr>
<td>6</td>
<td>13</td>
<td>0</td>
<td>2.0</td>
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<td>22</td>
<td>1</td>
<td>2.1</td>
</tr>
<tr>
<td>8</td>
<td>10</td>
<td>0</td>
<td>2.2</td>
</tr>
<tr>
<td>9</td>
<td>95</td>
<td>18</td>
<td>4.9</td>
</tr>
<tr>
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<td>0</td>
<td>2.0</td>
</tr>
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<td>11</td>
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<td>1.6</td>
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<td>39</td>
<td>7</td>
<td>4.6</td>
</tr>
<tr>
<td>13</td>
<td>18</td>
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<td>0.0</td>
</tr>
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<td>2</td>
<td>2.9</td>
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<td>2.9</td>
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<tr>
<td>19</td>
<td>63</td>
<td>9</td>
<td>5.6</td>
</tr>
</tbody>
</table>

Total: 643 101 (16%) 4.0
Similarly, CRIB score, gestation, and birthweight were all significant univariate predictors of neurological morbidity (P<0·0001). Neurological problems increased from 5% (22/450) with a CRIB score of 0–5 to 28% (5/18) with a score of >11. These problems increased from 2·7% (5/183) in those born at 31 weeks' gestation or more to 25% (2/8) in those born at less than 25 weeks' gestation, and increased from 3·5% (10/287) in those whose birthweight was >1249 g to 23% (6/26) in those weighing <750 g. However, for the prediction of morbidity, the ROC curve showed no difference between CRIB score, gestation, or birthweight (fig 2).

CRIB score, gestation at birth, and birthweight were correlated strongly with length of stay, the Spearman rank correlations being 0·612, −0·75, and −0·766, respectively.

The crude odds of death, the odds of death adjusted for CRIB score, and standardised mortality ratios are shown in table 3. After allowing for ex utero transfers, seven level 1 hospitals had zero mortality and were omitted from the table. There was no significant difference in mortality or adjusted mortality between the remaining 11 hospitals (P>0·5). The odds of death were higher, both before and after adjustment for case-mix in level 3, for high workload and designated neonatal units compared with the rest, but the differences were not significant (table 4).

New values for the separate components of CRIB, derived from the data in this study, gave increased weight to low gestation (<24 weeks). Low gestation scored 1/23 in the original CRIB and 6/32 in the new derivation. There was a small, significant improvement in fit when the score was based on four instead of two gestational age categories, but the change in the ROC curve of mortality in hospital was minimal.

**Discussion**

As far as we know, this is the first study to evaluate CRIB in a geographical cohort of infants. Data for the score were easy to collect and could readily be incorporated into any register of very low birthweight infants. To interpret the score, data from each centre need to be pooled, which means that individual hospitals lose sight of their own data and this reduces the impetus to collect it. The SMR method, which computes adjusted risk, can be calculated locally once expected mortality rates for each subgroup of the CRIB score are known. In this study the standard population, from which the expected mortality rates were derived, was the regional cohort. This is not necessarily the ideal standard and there is therefore a danger that hospitals will interpret a low SMR as a very good result when it may only be relatively good. Notwithstanding, table 3 shows that hospital ranking by SMR is almost identical with that using CRIB, the difference being due to small numbers and wide confidence intervals. Central collation of CRIB data could be justified by encouraging hospitals to use CRIB when calculating SMRs.

**Table 3** Comparison of hospital mortality rates and ranks

<table>
<thead>
<tr>
<th>Hospital*</th>
<th>Odds (No. of deaths)</th>
<th>Adjusted odds (rank)</th>
<th>SMR (rank)</th>
</tr>
</thead>
<tbody>
<tr>
<td>CRIB</td>
<td>Rank</td>
<td>CRIB</td>
<td>Rank</td>
</tr>
<tr>
<td>7</td>
<td>0·22 (1)</td>
<td>0·27</td>
<td>1</td>
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<tr>
<td>10</td>
<td>0·38 (4)</td>
<td>0·30</td>
<td>2</td>
</tr>
<tr>
<td>14</td>
<td>0·47 (2)</td>
<td>0·33</td>
<td>3</td>
</tr>
<tr>
<td>12</td>
<td>0·53 (7)</td>
<td>0·52</td>
<td>4</td>
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<tr>
<td>5</td>
<td>0·24 (1)</td>
<td>0·52</td>
<td>8</td>
</tr>
<tr>
<td>9</td>
<td>0·10 (18)</td>
<td>0·55</td>
<td>6</td>
</tr>
<tr>
<td>4</td>
<td>0·70 (6)</td>
<td>0·80</td>
<td>7</td>
</tr>
<tr>
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<td>9</td>
</tr>
<tr>
<td>3</td>
<td>1·34 (6)</td>
<td>1·20</td>
<td>10</td>
</tr>
<tr>
<td>11</td>
<td>2·35 (11)</td>
<td>1·48</td>
<td>11</td>
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</table>

*Seven hospitals had zero mortality. Although not excluded from the analysis, they are excluded from the table.
†Designated regional neonatal intensive care unit.
Compared with gestation and birthweight
CRIB was better at predicting mortality, was as
good at predicting morbidity, but was less
good at predicting length of stay. Date of
discharge from hospital is an arbitrary end-
point, influenced by differing policies of care
rather than quality of care or illness severity.
However, it corresponds well to post-concep-
tional age and was remarkably constant for
babies of all gestational ages in this study (fig
3). Thus it seems reasonable to use mortality
in hospital as an outcome measure. CRIB was
then the best predictor of mortality in hospital.
At a false positive rate of 5%, CRIB had a
sensitivity of 67% compared with a sensitivity
of 41% for gestation and for birthweight.

A score which can accurately predict mortal-
ity could be used by purchasers of health care
to refuse to provide expensive care for babies
with a high score. This would be an abuse of an
epidemiological tool in individual cases. There
were 16 babies with a CRIB score of 16 or
more who performed their own triage. Fifteen
died after using less than a total of 25 days of
intensive care, a fraction of the per cent of days
spent in category 1 intensive care each year.
Though the decision to treat very small babies
poses many ethical problems, the CRIB score
demonstrates that cost is a minor factor.

Cranial ultrasound findings in the very low
birthweight baby have been shown to correlate
strongly with their neurodevelopmental out-
come at 4 years of age and CRIB has been
shown to be highly correlated with brain ultra-
sound findings. It does not necessarily follow
that CRIB can predict neurodevelopmental
outcome. In the current study, although CRIB,
gestation at birth, and birthweight were all
significantly associated with cranial ultrasound
findings, they all had low sensitivity of about
20% at a 5% false positive rate. Thus the
CRIB score as currently calculated does not
improve on existing data (such as birthweight
and gestational age) for the prediction of brain
ultrasound abnormalities and, unlike gestation
at birth and birthweight, babies have not yet
been followed up for long enough to demon-
strate an association with neurodevelopmental
outcome.

Traditionally, birthweight specific mortality
rates have been used to assess the outcome of
neonatal and obstetric care but despite evidence
that gestation at birth is a superior measure.

The ROC curve of mortality in hospital also
shows that gestation at birth is better than
birthweight. It is surprising, therefore, that ges-
tation at birth made such a small contribution
to the overall CRIB score, and it has been
excluded in other similar scores. Yet, more
emphasis on gestation in the CRIB score failed
to improve its application in practice.

Hospitals were defined as teaching or non-
teaching in the original paper on CRIB depending on where medical students received
most of their clinical training. This definition
gave no clear idea as to the real differences
among hospitals with regard to neonatal care
service provision. Classifications by workload
or level of care are reproducible and compar-
able, and future studies on hospital perform-
ance should use similar standardised criteria.
The interpretation of CRIB when used to mea-
sure hospital performance was difficult. Even
over a one year period seven hospitals had too
few cases to create a meaningful score and, due
to the requirement for the implementation of a policy of intrauterine transfer of high risk
infants, they had mortality rates of zero. If
CRIB was functioning as a measure of perfor-
mance the ‘best’ hospitals (level 3, high work-
load, designated NICU) would be expected to
have a higher rank. This did not occur either
because the level 3 hospitals in South East
Thames were performing badly or because CRIB was not sensitive to the performance it
permitted to measure. The excess mortality
may equate with unmeasured risk rather than
ineffective care. Furthermore, mortality may
not be the best measure of quality of care as
hospitals with low mortality may discharge
babies with higher levels of morbidity. Until
CRIB is evaluated further it cannot yet be used
by purchasers or providers of neonatal
intensive care as a sensitive indicator of
hospital performance.

Thanks are due to South East Thames Regional Health
Authority for funding (R de C-W, MS) and to the health care
professionals in all the units for permitting access to their
data.

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Figure 3 Median gestational age at discharge from
hospital (interquartile range) in survivors (n = 542).


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