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MRI-based brain volumes of preterm infants at term: a systematic review and meta-analysis

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► Additional supplemental material is published online only. To view, please visit the journal online (<http://dx.doi.org/10.1136/archdischild-2021-322846>).

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Received 22 July 2021

Accepted 30 November 2021

Published Online First

25 January 2022

ABSTRACT

Background MRI allows a detailed assessment of brain structures in preterm infants, outperforming cranial ultrasound. Neonatal MR-based brain volumes of preterm infants could serve as objective, quantitative and reproducible surrogate parameters of early brain development. To date, there are no reference values for preterm infants' brain volumes at term-equivalent age.

Objective Systematic review of the literature to determine reference ranges for MRI-based brain volumes of very preterm infants at term-equivalent age.

Methods PubMed Database was searched on 6 April 2020 for studies reporting MR-based brain volumes on representative unselected populations of very preterm and/or very low birthweight infants examined at term equivalent age (defined as 37–42 weeks mean postmenstrual age at MRI). Analyses were limited to volumetric parameters reported in >3 studies. Weighted mean volumes and SD were both calculated and simulated for each parameter.

Results An initial 367 publications were identified. Following application of exclusion criteria, 13 studies from eight countries were included for analysis, yielding four parameters. Weighted mean total brain volume was 379 mL (SD 72 mL; based on n=756). Cerebellar volume was 21 mL (6 mL; n=791), cortical grey matter volume 140 mL (47 mL; n=572) and weighted mean volume of unmyelinated white matter was 195 mL (38 mL; n=499).

Conclusion This meta-analysis reports pooled data on several brain and cerebellar volumes which can serve as reference for future studies assessing MR-based volumetric parameters as a surrogate outcome for neurodevelopment and for the interpretation of individual or cohort MRI-based volumetric findings.

INTRODUCTION

Preterm birth is a major contributor to the global burden of disease.¹ Neurodevelopmental outcome after preterm birth ultimately determines quality of life; however, the determinants of this outcome are still subject to research.^{2–5}

Impaired neurodevelopmental outcome may be present without overt intracerebral lesions detectable by neonatal cranial ultrasound.⁶ In order to identify valid biomarkers that facilitate the prediction of neurocognitive long-term outcome and, at the same time, establish early surrogate endpoints for clinical studies aimed at improving perinatal care, various parameters have been studied on cerebral MRI in preterm infants in the last decades. Of these, advanced methods yielding quantitative parameters (such as volumetric MRI, diffusion

What is already known on this topic?

- ⇒ Very preterm and very low birthweight infants show smaller brain volumes compared with term born infants.
- ⇒ Smaller cerebral volumes in preterm infants are associated with impaired neurocognitive outcome.
- ⇒ No systematic reference data are available to date regarding brain volumes of very preterm/very low birthweight infants examined at term.

What this study adds?

- ⇒ A systematic search for studies reporting MR-based brain volumes on preterm infants at term equivalent age was performed and weighted means were calculated.
- ⇒ These could serve as reference data for future studies examining cerebral volumes in preterm infants at term.

MRI, ¹H-MR spectroscopy and resting-state functional connectivity) seem to be most promising for providing objective, predictive and sensitive values.⁷

This systematic review focuses on volumetric MRI, aiming at identifying reference ranges for cerebral and cerebellar volumes in unselected very preterm/very low birthweight (VLBW) infants at term-equivalent age. There is evidence suggesting that reduced brain tissue volumes in preterm infants are associated with impaired neurocognitive outcomes.^{8–13} Reduced cerebral and cerebellar tissue volumes may result from loss of immature oligodendrocytes which are vulnerable to inflammation and fluctuations in haemodynamics and oxygenation eventually resulting in altered cerebral development and impaired function.^{8 14–17}

METHODS

To identify all published studies reporting brain volume measurements of very preterm or VLBW infants, PubMed was searched on 6 April 2020 using the following systematic search strategy: ('infant, premature' (MeSH) OR 'infant, low birth weight' (MeSH)) AND 'magnetic resonance imaging' (MeSH) AND (volum*(All Fields)). The term 'brain' was intentionally not included in the search strategy to avoid omitting studies focusing on, for example, cerebellar/hippocampal volume



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To cite: Romberg J, Wilke M, Allgaier C, et al. *Arch Dis Child Fetal Neonatal Ed* 2022;**107**:F520–F526.

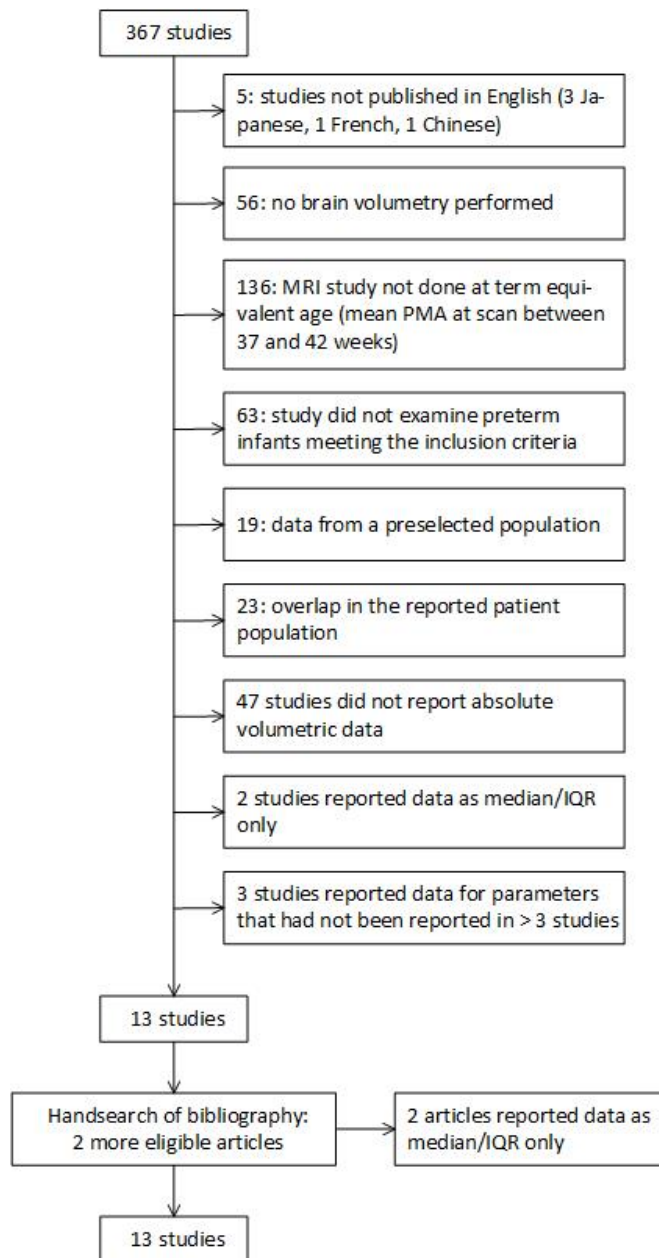


Figure 1 Flow diagram for systematic review of retrieved studies. PMA, postmenstrual age.

measurements. The process of reviewing and assessing studies was based on the guidelines established by Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA).

Articles were selected for data extraction if they fulfilled all of the following criteria:

- ▶ Reporting volumetric data of the brain or its anatomical structures acquired by structural MRI.
- ▶ Examining very preterm infants (ie, born at a gestational age <32 weeks) or VLBW infants (ie, with birth weight <1500 g).
- ▶ Reporting MRI data acquired at term-equivalent age (ie, if the mean postmenstrual age (PMA) at scan was between 37 and 42 weeks).
- ▶ Including representative populations (ie, studies reporting preselected subpopulations were excluded).
- ▶ Published in English.

- ▶ Accessible (either online, via interlibrary loan, or by directly contacting the authors).

Further,

- ▶ If several manuscripts reported data from the same population, only the study including the largest number of infants was included in this review. If there was only a small overlap between patient populations (<25%), both studies were included.
- ▶ Only brain imaging outcome parameters reported in >3 studies were included in the meta-analysis.

First, title and abstract of all articles identified with the search strategy were reviewed manually (by JR and CA) with regard to the above-mentioned inclusion and exclusion criteria, and studies not satisfying these criteria removed. In potentially eligible studies, the full text of each article was reviewed. The bibliography of each eligible full-text article was hand-searched for additional studies. If studies reported concurrent populations of term-born infants, brain volume data were extracted and reported, too.

The following quality criteria were assessed and reported for all eligible studies: Were the infants representative of the underlying population? Was the reliability of image processing methods tested?

Cohorts of included studies were assessed regarding the following characteristics: criteria for inclusion and exclusion in the MRI study, site and time period of recruitment, mean birth weight, mean gestational age at birth, proportion of small for gestational age (SGA) infants, use of antenatal and postnatal steroids, proportion of infants with bronchopulmonary dysplasia (BPD), persistent ductus arteriosus (PDA) and intraventricular haemorrhage (IVH), postmenstrual age (PMA) at MRI scan and type of MRI scanner.

Statistical analyses

All volumes are reported in millilitres. Gestational age at birth and PMA at MR scan are reported in weeks. In case data had been reported by subgroups, the weighted mean for the overall study population was calculated based on all subgroups.

Meta-analysis of volumetric data was performed in two ways:

(A) Weighted mean values and weighted SD were calculated, the latter on the basis of total variance using SAS statistical software version 9.4:

$$\text{Total variance: } s_{\text{total}}^2 = \frac{1}{n} \sum_{i=1}^M n_i * s_i^2 + \frac{1}{n} \sum_{i=1}^N w_i * (x_i - x^*)^2$$

(B) Data from all studies for a given parameter were simulated using the Matlab random number generator (The MathWorks, Natick, Massachusetts, USA): Based on the mean (SD) of a particular study, a total of n populations with normally distributed 100 000 data points each were generated, with n corresponding to the number of subjects included in the original studies. These synthetic populations of all studies were combined by means of histogram analyses to yield a final population. To this new population (representing all contributing populations in accordance with their original size), we fitted a Gaussian distribution from which a final population mean and SD, as well as a goodness-of-fit parameter was reported.

Linear regression for brain volumes in relation to mean age at scan was calculated and visualised using Python V.3.7 (<https://www.python.org/>).

Registration

This review was not registered. A review protocol was prepared, but not published.

Table 1 Demographic variables of study populations

Reference	N	Inclusion criteria	Site of recruitment	Study period	Birth weight (g) Mean (SD)	Gestational age (weeks) Mean (SD)	SDS (birth weight) (Mean(SD))/ [§] % IUGR*	% Antenatal Corticosteroids (any)	% Postnatal Corticosteroids	% BPD	% PDA	% any IVH	PMA at scan (weeks) Mean (SD)	MR scanner	Segmentation method
Blok <i>et al</i> , PLoS One, 2014 ²⁰	36	GA <32 weeks	Utrecht (NL)	2007–2009	1027 (47)	27.2 (1.7)	N/A	N/A	N/A	42%	28%	%I-V: 38.9%, %II: 14%, %III: 5.6%, %IV: 11%, %V: 8.3%	1.5 T Philips Intera	Semiautomatic	
Ehlsjeld <i>et al</i> , J Pediatr, 2010 ³⁰	209	bw<1500 g and GA <32 weeks	Turku (FI)	2001–2006	1121 (319)†	29.0 (2.7)	-1.43 (1.61)/38% (SGA)†	N/A	14%†	14%†	12.5% (missing information for 4)	%I-V: 23.7%, †, †	0.23 T Outlook GP Philips (n=126)/1.5 T Philips Gyroscan Intera (n=106) (23 scans excluded)	Manual	
Hansen-Popper <i>et al</i> , J Clin Endocrinol Metab, 2011 ²²	51	GA <31 weeks	Lund (SE)	2005–2007	888 (288)	26.4 (1.9)	24%	N/A	N/A	N/A	N/A	§	3 T Siemens Magnetom Allegra	Manual	
Hansen-Popper <i>et al</i> , Pediatr Res, 2013 ³¹	49	GA <31 weeks	Lund (SE)	2005–2007	889 (290)	26.0 (1.9)	-1.0 (1.2)	N/A	N/A	N/A	N/A	N/A	3 T Siemens Magnetom Allegra	Manual	
Inder <i>et al</i> , Pediatrics, 2005 ⁸	119 (of 129)**	bw<1500 g and GA ≤32 weeks	n=100: Christchurch (NZ), n=29: Melbourne (AU)	1998–2000	1040 (272)	27.7 (2.2)	22%	87%	2%	21%	32%	%I-V: 17%	1.5 T General Electric Signa System	Semiautomatic	
Kamino <i>et al</i> , Pediatr Res, 2018 ³⁸	44 (of 60)††	GA <32 weeks	n=33: San Francisco (USA), n=7: Vancouver (CA)	2010–2011	N/A	UCSF: 28.1 (1.8) UBC: 28.5 (2)	N/A	UCSF: 100% UBC: 63%	N/A	N/A	UCSF: 48% UBC: 37%	%I: 0%, UCSF: 12%, UBC: 19% %II–IV: UCSF: 1.8%, UBC: 4%	1.5 T General Electric Signa/Siemens Avanto	Automatic	
Moeskops <i>et al</i> , PLoS One, 2015 ²¹	85	GA <28 weeks	Utrecht (NL)	2008–2013	N/A	26.6±1.0	N/A	N/A	N/A	N/A	N/A	%I: 11.8%, %II: 15.3%, %III: 5.9%	3 T Philips Achieva	Automatic	
Parikh <i>et al</i> , PLoS One, 2013 ³⁹	122	bw<1000 g	Houston (USA)	2005–2007	771 (144)	26.7 (1.9)	25% (SGA)	70%	18% postnatal steroids for BPP	52%	33% with PDA surgery	%I-V: 18%	1.5 T Reliant Medical GE-LX	Semiautomatic	
Steinhom <i>et al</i> , J Pediatr, 2015 ⁴⁴	192	bw<1250 g and/or GA <30 weeks	Melbourne (AU)	2001–2003	956 (223)	27.5 (1.9)	9%	89%	9%	N/A	N/A	%II-IV: 4%	1.5 T General Electric Signa System	Manual	
Stiver <i>et al</i> , J Child Neurol, 2015 ⁴⁵	105	GA <32 weeks	Toronto (CA)	2008–2010	N/A	28.8 (1.8)	N/A	69%	6% postnatal steroid exposure†	N/A	N/A	%III: 1.9%, %II/IV: 21%	1.5 T Signa Twin EXCITE3	Semiautomatic	
Tam <i>et al</i> , Sci Transl Med, 2011 ⁴⁷	68 (of 172)¶	GA<33 weeks	San Francisco (USA)	2006–2009	N/A	27.5 (2.1)	N/A	90%	22%	51%	50%	%I-V: 32.4%	1.5 T General Electric Signa/Siemens Avanto	Semiautomatic	
Thompson <i>et al</i> , Brain, 2007 ²⁵	202	bw<1250 g and/or GA<30 weeks	Melbourne (AU)	2001–2004	965 (239)	27.6 (2.0)	11%	87%	9%	34%	36% (Indomethacin administered)	%I-V: 12%	1.5 T General Electric Signa System	Manual	
Vasur <i>et al</i> , BMJ Open, 2014 ³¹	19 (of 22)**	GA<32 weeks	London (UK)	2007–2008	1260 (410)	28.9 (2.8)	-0.13 (0.78)	N/A	N/A	25%	20% with pharmacological therapy	%II-IV: 5%	3 T Philips Achieva	Automatic	

N/A, data not available in the published manuscript. Brain volumetry was not available in all infants. *Proportion with SDS for birth weight <2SD marked as 'SGA', then proportion with birth weight <10th percentile. †According to publication: Manu *et al*, PIPARI Study Group. Ventricular dilatation in relation to outcome at 2 years of age in very preterm infants: a prospective Finnish cohort study. Dev Med Child Neurol. 2011. Jan;53(1):48–54. ‡According to personal communication with the author: MRI were performed at term (SD 5 days) in 93% of infants. The time between term age and imaging day was at most 29 days (in one infant). §Only the percentage of 'severe brain damage'; I-V, III and/or parenchymal haemorrhage and/or WMD was reported (12%). ¶From the original data reported by Tam *et al*, the infants meeting the inclusion criteria of this review were selected and analysed. **Only the percentage of infants with hydrocortisone treatment was reported. ††Patient characteristics were reported from the original population

bw, birth weight; GA, gestational age at birth; IVH, intraventricular haemorrhage; SGA, small for gestational age; UCSF, University of California San Francisco; WMD, white matter disease.

Table 2 Brain volume measurements (and weighted means) in preterm infants at term-equivalent age

	N	PMA at scan (weeks) Mean (SD)	Total brain volume (mL) Mean (SD)	Total cerebellar volume (mL) mean (SD)	Cortical grey matter volume (mL) Mean (SD)	Unmyelinated white matter volume (mL) Mean (SD)
Blok <i>et al</i> , 2014 ²⁰	36	41.4 (0.8)	391 (34)	N/A	167 (23)* (overlap with Moeskops <i>et al</i> ²¹)	161 (21)* (overlap with Moeskops <i>et al</i> ²¹)
Ekblad <i>et al</i> , 2010 ³⁰	209	†	397 (51.2)	24.2 (5.1)	N/A	N/A
Hansen-Pupp <i>et al</i> , 2011 ²²	51	40.1 (0.6)	396 (46.3)* (n=46, overlap with Hansen-Pupp <i>et al</i> ²³)	21.3 (2.7)	N/A	191 (25)* (n=46, overlap with Hansen-Pupp <i>et al</i> ²³)
Hansen-Pupp <i>et al</i> , 2013 ²³	49	40.1 (0.6)	393 (50.9)	21.1 (3.5)* (overlap with Hansen-Pupp <i>et al</i> ²²)	N/A	190 (25)
Inder <i>et al</i> , 2005 ⁸	119	40.2 (0.3)	406 (57)	N/A	178 (41)	202 (41)
Kamino <i>et al</i> , 2018 ²⁸	44	37.7 (3.0)	N/A	16.7 (5.3)	107.6 (40.2)	151 (20.1)
Moeskops <i>et al</i> , 2015 ²¹	85	41.1 (0.5)	N/A	N/A	107 (13)	171 (19)
Parikh <i>et al</i> , 2013 ²⁹	122	38.5 (2.2)	270 (41.5)	16.1 (4.2)	105 (19.1)	N/A
Steinhorn <i>et al</i> , 2015 ²⁴	192	‡	395 (63.2)* (overlap with Thompson <i>et al</i> ²⁵)	21.4 (4.5)	159 (41)* (overlap with Thompson <i>et al</i> ²⁵)	N/A
Stiver <i>et al</i> , 2015 ²⁶	105	41.9 (2.0)	N/A	23.9 (5)	N/A	N/A
Tam <i>et al</i> , 2011 ²⁷	68	39.5 (1.3)	N/A	20.4 (5)	N/A	N/A
Thompson <i>et al</i> , 2007 ²⁵	202	40.1 (1.7)	395 (64)	N/A	159 (41)	212 (32)
Vasu <i>et al</i> , 2014 ³¹	19	40.2 (3.1)	462 (12.8)	N/A	N/A	N/A
Aggregate data all studies Weighted mean ($\sqrt{\text{Total variance}}$)			379 (72); n=756	21 (6); n=791	140 (47); n=572	195 (38); n=499
Aggregate data all studies Mean (SD) by simulation			391 (67)§	19 (5)§	135 (68)§	192 (38)§
Aggregate data excluding reference ²⁹ Weighted mean ($\sqrt{\text{Total variance}}$)			399 (57) n=634	22.3 (5.3) n=669	149.2 (47.8) n=450	N/A

*Not included in the meta-analysis.
†According to personal communication with the author: 'MRI were performed at term (SD 5 days) in 93% of infants. The time between term age and imaging day was at most 29 days (in one infant).' Out of 232 total MRI scans, 209 were successfully performed and analysed.
‡Between 38 and 42 weeks.
§Adjusted R² for TBV 0.9850, for CV 0.9965, for CGM 0.8822, for UMWM 0.9906.
CGM, cortical grey matter; CV, cerebellar volume; N/A, not available; PMA, postmenstrual age; TBV, total brain volume; UMWM, unmyelinated white matter.

RESULTS

Using our initial search criteria, a total of 367 studies, published between 1992 and 2019, were retrieved. After review for inclusion and exclusion criteria, 13 were included (see [figure 1](#)). Hand-searching the bibliographies yielded two additional studies meeting the inclusion criteria.^{18 19} However, both could not be included because volume data were reported as median (IQR) and least-squares means, respectively. This left 13 articles for inclusion in the meta-analysis.

We identified potential overlap in populations between studies.²⁰⁻²⁵ In these cases, for each volumetric parameter, only the study with the largest population was included for the respective analysis. A small overlap in the population of two studies^{8 25} resulted in 29 infants being included twice.

Characteristics of the included study populations are listed in [table 1](#). Studies indicated the following reasons for exclusion of subjects: Major congenital anomalies (stated in 9 out of 13 studies) and low-quality MR images (stated in 5 out of 13 studies). All studies excluded infants with missing parental consent. Three studies additionally excluded infants with congenital infection (no further definition given).²⁶⁻²⁸

Only one study tested representativeness of its study population, and there were no differences between included and excluded subjects regarding gender, multiple birth, gestational age, IVH and BPD.²⁵ Some studies indicated recruitment rates in relation to eligible subjects, which varied considerably

(90%/80%,⁸ 73%,²¹ 100%,²⁹ 66%,²⁴ 94%³⁰). Most studies (10/13) were single-centre cohorts from level III academic centres in highly developed countries.

All studies reported MRI postprocessing and segmentation methods. Except for one study,³¹ all studies specified reliability testing for their technique. Studies employed manual, semiautomated and in three studies,^{20 21 29} automated segmentation techniques (see [table 1](#)).

Brain volume measurements

The following volumetric parameters were reported in >3 studies: total brain volume (TBV; n=756), cerebellar volume (CV; n=791), volume of cortical grey matter (CGM; n=572) and volume of unmyelinated white matter (UMWM; n=499). Brain volumes are presented in [table 2](#). [Figure 2](#) displays the relation between volume measurements and mean PMA at scan. Simulated populations and fitted Gaussian distribution are depicted in [figure 3](#).

Four studies^{8 25 29 31} also reported MRI-based brain volumetry in contemporary samples of term-born infants. Results are summarised in online supplemental table S1 and indicate that CGM and UMWM volumes might be reduced by about 25% in preterm compared with term infants.

Based on the simulated population data, an online supplemental MS-excel spread sheet is provided converting MRI brain volumes of preterm infants into z-scores.

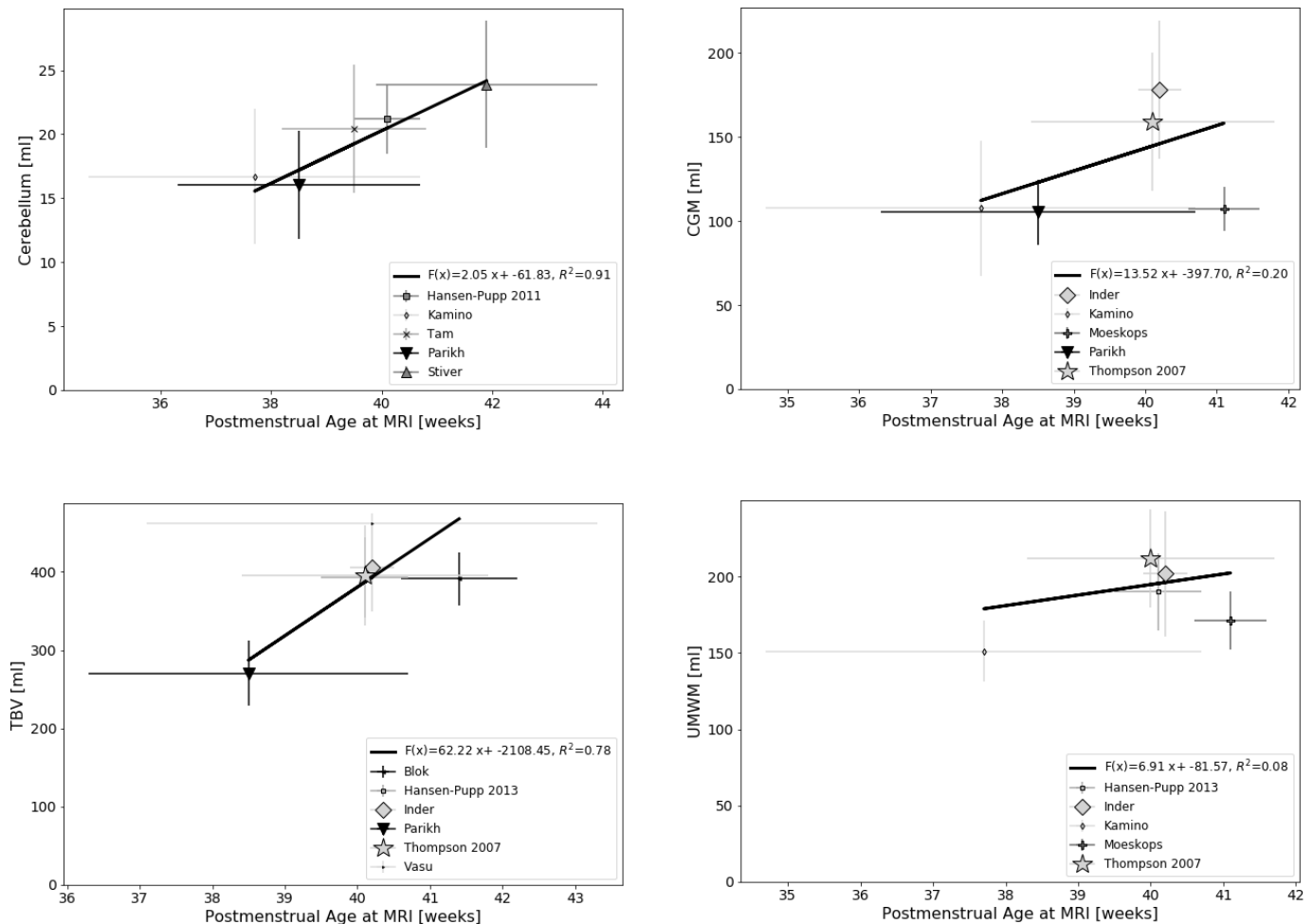


Figure 2 Brain volumes by postmenstrual age (PMA) at MRI scan. CV, cerebellar volume; CGM, cortical grey matter volume; TBV, total brain volume; UMWM, unmyelinated white matter volume. Panels showing mean cerebral volumes according to mean PMA at MRI scan (with horizontal error bars indicating SD for PMA at MRI scan and vertical error bars indicating SD for volume). Sample size for each study is represented by the size of the symbol. Linear regression lines were calculated for each parameter taking sample size into account.

DISCUSSION

For this systematic review, we identified 13 studies reporting MRI-based brain volume measurements in very preterm infants at term-equivalent age. For TBV, CV, CGM and UMWM, sufficient data were available for this meta-analysis.

All studies included infants with neonatal complications such as BPD, PDA and IVH. Since exclusion criteria of the studies were limited to missing parental consent, 'congenital anomalies' and low-quality MRI images and recruitment rates were $\geq 67\%$ (reported in six studies), we can assume these study populations to be representative of very preterm infants. However, most studies included single-centre cohorts from academic level III centres in highly developed countries, limiting generalisability to some extent.

In one study, inclusion was restricted to infants ≤ 1000 g birth weight²⁹; consequently, mean birth weight was lower than in the other study populations. Mean TBV, CV and CGM volume all were considerably lower in this cohort compared with the other preterm populations. This discrepancy can partly be attributed to the fact that the timepoint of MRI measurements in this study (mean PMA at scan: 38.5 ± 2.2 weeks) was rather early compared with the other studies. Another explanation might be that ELBW infants (i.e. birthweight < 1000 g) are more prone to impaired brain growth than VLBW infants; these hypotheses

should be tested in future studies or individual patient data meta-analyses. A number of studies (including that of Parikh *et al*²⁹) found smaller brain volumes in IUGR/SGA infants.^{8 22 25 29 32 33} Notably, Parikh *et al* also found brain volumes of contemporary term born infants to be markedly smaller compared with those reported by other studies. Thus, a systematic measurement bias could also have contributed to the observed differences. For reference purposes, we also provide the weighted mean and SD values excluding Parikh's study for TBV, CV and CGM volume.

Figure 2 visualises the variance of measured brain volumes among the studies, highlighting the need of large study groups and pooled data. Additionally, linear regression was calculated for brain volumes by PMA at scan showing the best fit for a linear relation for CV. Measurements of CV could be more reliable due to clear anatomical boundaries, resulting in the lower variability. Furthermore, studies reporting CV were well dispersed for PMA at scan. According to Gui *et al*,³⁴ the cerebellum shows one of the highest growth rates around term, which could explain why a growth trend can be illustrated here. For the other parameters, variance in this restricted time frame is too high to deduce growth trends from the data.

This systematic review and meta-analysis has several limitations: Only one database (PubMed) was searched and only articles in English were included. However, only a very small number

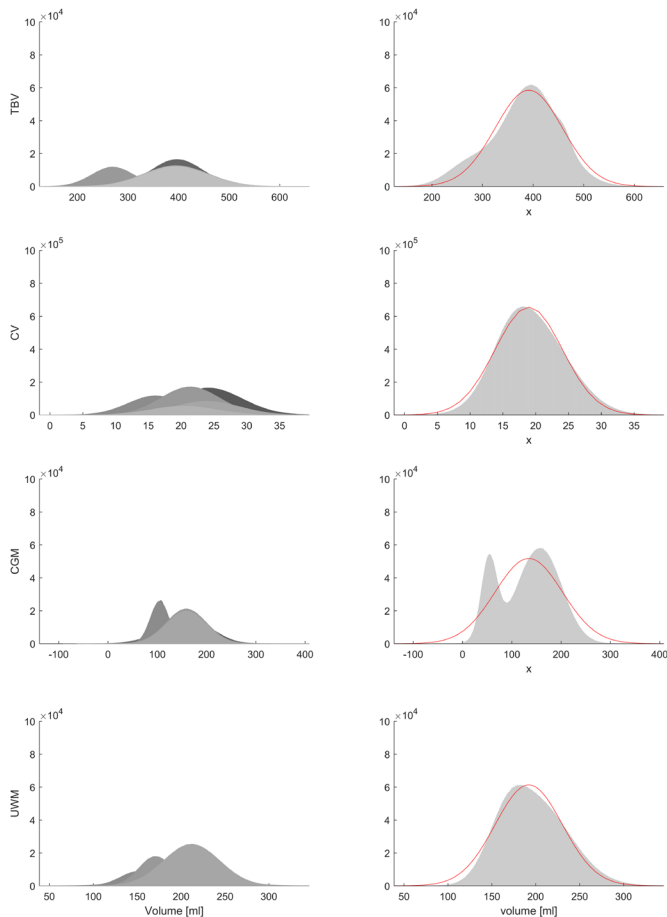


Figure 3 Meta-analysis by simulation studies. CV, cerebellar volume; CGM, cortical grey matter volume; TBV, total brain volume; UMWM, unmyelinated white matter volume. Panels in the left column depict individual synthetic populations (according to simulations described in the Methods section), right column depicts summarising histograms and fitted Gaussian distributions. Resulting means, SDs and R^2 s are reported in table 2. Figure with permission and created by MW.

(5/367) of initial articles were not published in English and the likelihood that any neonatal volumetric MRI study would not be indexed in PubMed was considered low. Hand-searching the bibliography of all eligible articles also reduced the probability of missing relevant studies.

The number of included studies was relatively low¹³ compared with the number of initially retrieved studies (367). This is due to a combination of relatively strict inclusion criteria and a broad search strategy which retrieved numerous studies in older subjects. Furthermore, studies deviating from our definition of 'term equivalent age' (mean PMA at scan between 37 and 42 weeks) were excluded.

Studies excluding infants with brain lesions (eg, Padilla *et al*³⁵) were not included since our intention was to report on representative preterm infant cohorts. In contrast to our study, Padilla *et al*³⁵ showed higher mean values in their cohort regarding CGM (171,2 mL) and the CV (26,6 mL). This is consistent with the findings of other studies.^{36–38}

Another limitation is that included cohorts ranged from 1998 to 2013, and by our predefined criteria no more recent cohorts were selected, possibly causing bias since both imaging and clinical treatment have evolved. Among the included studies, however, no secular trend could be seen. Furthermore,

considerable variability in image acquisition and processing was recognised. For brain volume extraction, tissue segmentation into grey and white matter and cerebrospinal fluid (CSF) is the crucial step.³⁹ Of the included studies, five used a manual, five a semiautomatic and three an automatic approach (cf. table 1). All but one study³¹ reported internal validation of their segmentation process. While a congruent approach would of course have been preferable, this (incidentally rather well-balanced) composition makes a substantial bias resulting from our data processing approach unlikely.

Despite heterogeneity concerning study procedures and reporting practices, relatively congruent brain volumes were reported, apart from the ELBW study.²⁹ In total, a substantial number of infants could be included in the meta-analysis. It is also reassuring that both methods used to combine the values across populations, despite taking very different approaches, also yielded very similar values (cf. table 2). Thus, the resulting mean values and SDs presented in this review can serve as reference data for future volumetric MRI studies among preterm infants and can be consulted when interpreting individual or cohort brain volume measurements in VLBW/very preterm infants. As online supplemental data, we provide a calculator embedded within an Excel worksheet for easy comparison of individual versus these group values.

Ultimately, brain volumes of term born infants should serve as reference for regular, intrauterine brain growth. However, despite all efforts, there is still a gap between cohorts of term born and very preterm infants. In 4 of 13 studies included in this review, concurrent cohorts of term infants were scanned around term-equivalent age, and significantly smaller total brain tissue volumes were demonstrated for preterm infants in all four studies.^{8 25 29 31} According to Thompson *et al*²⁵ and Inder *et al*,⁸ this reduction is driven by smaller CGM and DGM volumes. Furthermore, all four studies revealed significantly increased total CSF volumes in preterm infants, likely reflecting loss or inadequate growth of cerebral tissue (ie, grey or white matter).

Although brain growth of term born infants should be the ultimate goal, brain volume data of unselected very preterms are required when planning interventional studies aiming at optimising brain growth on a realistic basis.

As mentioned above, several studies found an association between birth weight and brain volumes at term. Additionally, male sex^{25 32 40} and higher gestational age at birth^{8 22} were found to be associated with larger brain volumes. While it would have been interesting to separate out brain volumes with regard to these factors, the respective data were not available in sufficient detail to allow for their inclusion in multivariate analyses. Future studies on MRI brain volumes in VLBW/very preterm infants should aim at further identifying relevant factors that influence early brain volume growth after preterm birth and, if possible, at quantifying their impact. In this context it would be of special interest to address those factors that are not predetermined (such as sex) but can be modified by optimised neonatal intensive care (such as nutrition, oxygenation, etc). Further studies are also needed that investigate the long-term predictive value of individual MR-based brain volume measurements at term with respect to neurocognitive outcome in childhood and beyond.

In conclusion, in this meta-analysis of 13 studies including data from more than 900 very preterm infants, with the exception of one study,²⁹ rather similar results for TBV, CV and UMWM volumes were found at term-equivalent age. These weighted mean volumes may serve as reference for individual patient data and future studies.

Contributors AF, MW and JR conceived the idea of this systematic review. AF and JR developed the research strategy and criteria of inclusion and exclusion. JR and CA performed the research and selection of studies. AF, MW and JR devised the concept of data extraction, calculation and presentation. MW, CA, CE and JR performed the calculations. JR wrote the manuscript with support of MW, CFP, AF, CA and TN, who all made important intellectual contributions and approved the final version of the manuscript. MW and AF planned and designed the supplemental material. JR acts as guarantor for the study.

Funding The authors have not declared a specific grant for this research from any funding agency in the public, commercial or not-for-profit sectors.

Competing interests None declared.

Patient consent for publication Not applicable.

Ethics approval This study does not involve human participants.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement All data relevant to the study are included in the article or uploaded as supplementary information. The authors confirm that the data supporting the findings of this study are available within the article and/or its supplementary materials.

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