

Gestational age and hospital admission costs from birth to childhood: a population-based record linkage study in England

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► Additional supplemental material is published online only. To view, please visit the journal online (http://dx.doi. org/10.1136/archdischild-2022-324763).

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Received 9 August 2022 Accepted 27 January 2023 Published Online First 9 February 2023

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Objective To examine the association between gestational age at birth and hospital admission costs from birth to 8 years of age.

ABSTRACT

Design Population-based, record linkage, cohort study in England.

Setting National Health Service (NHS) hospitals in England, UK.

Participants 1018136 live, singleton births in NHS hospitals in England between 1 January 2005 and 31 December 2006.

Main outcome measures Hospital admission costs from birth to age 8 years, estimated by gestational age at birth (<28, 28–29, 30–31, 32, 33, 34, 35, 36, 37, 38, 39, 40, 41 and 42 weeks).

Results Both birth admission and subsequent admission hospital costs decreased with increasing gestational age at birth. Differences in hospital admission costs between gestational age groups diminished with increasing age, particularly after the first 2 years following birth. Children born extremely preterm (<28 weeks) and very preterm (28-31 weeks) still had higher average hospital admission costs (£699 (95% CI £419 to £919) for <28 weeks: £434 (95% CI £305 to £563) for 28–31 weeks) during the eighth year of life compared with children born at 40 weeks (£109, 95% CI £104 to £114). Children born extremely preterm had the highest 8-year cumulative hospital admission costs per child (£80559 (95% CI £79238 to £82019)), a large proportion of which was incurred during the first year after birth (£71997 (95% CI £70866 to £73097)). **Conclusions** The association between gestational age at birth and hospital admission costs persists into mid-childhood. The study results provide a useful costing resource for future economic evaluations focusing on preventive and treatment strategies for babies born preterm.

BACKGROUND

The rates of preterm birth (<37 weeks' gestation) have increased or remained stable over the past few decades in most countries, ^{1–3} accounting for 10.6% of all live births worldwide in 2014.³ Survival rates following preterm birth have increased as a result of technological advances,^{4 5} but these babies still remain at a higher risk of infant mortality and a range of short-term and long-term morbidities.^{6 7} A recent study examined the association between

WHAT IS ALREADY KNOWN ON THIS TOPIC

- ⇒ Most previous research on the economic consequences associated with gestational age focused on the costs of initial birth admission or costs incurred during the first few years after birth.
- ⇒ A small number of studies have examined the association between gestational age at birth and hospital costs over the longer term, but were based on regional data, decision-analytic models synthesising summary evidence from multiple sources, cross-sectional assessments at specific ages, or focused on narrow categories of the gestational age range.

WHAT THIS STUDY ADDS

⇒ Using a large national cohort with hospital records linked from birth until mid-childhood, our study quantifies hospital admission costs from birth up to 8 years of age across the full range of gestational age at birth.

HOW THIS STUDY MIGHT AFFECT RESEARCH, PRACTICE OR POLICY

⇒ The results should act as a useful resource for clinical and budgetary service planning, and as data inputs for economic evaluations of preventive and treatment strategies for babies born at different gestational ages.

gestational age at birth and hospital admissions and found that gestational age at birth is a strong predictor of severe morbidity throughout childhood, even for those born at 38 and 39 weeks' gestation.⁸

Most previous research on the economic consequences associated with gestational age focused on the costs of initial birth admission or costs incurred during the first few years after birth.⁹ A small number of studies have investigated hospital costs by gestational age over the longer term, but were based on historical region-specific data,¹⁰ decisionanalytic models synthesising summary evidence from multiple sources,¹¹ ¹² cross-sectional assessments at a specific age,¹³ or focused on narrow categories of the gestational age spectrum.¹⁴ To the best of our knowledge, no study has comprehensively estimated the long-term economic burden



associated with gestational age from birth to middle childhood across the full spectrum of gestational age using a large national cohort of children born in the 21st century.

In this study, we conducted an evidence synthesis exercise to examine the association between gestational age at birth and hospital admission costs from birth to 8 years of age using a population-based, record-linkage study that included all live, singleton births occurring in England in 2005/2006 as part of the TIGAR study (Tracking the Impact of Gestational Age on Health, Educational and Economic outcomes: a Longitudinal Records Linkage Study).⁸

METHODS

Data sources

In this study, we synthesised data from three data sources: TIGAR, National Neonatal Research Database (NNRD) and Paediatric Intensive Care Audit Network (PICANet) databases. Access to individual patient-level data was available for the TIGAR dataset, whereas aggregate data were available from the NNRD and PICANet datasets.^{15 16}

The TIGAR cohort was built through a population-based data linkage using data from the Office for National Statistics (ONS) birth registration records linked to death registration records, birth notification records and Hospital Episode Statistics Admitted Patient Care (HES APC) records.¹⁷ A description of the datasets, linkage and quality assurance has been published elsewhere.^{8 18 19} In brief, the TIGAR cohort included all live, singleton births occurring in an NHS hospital in England between 1 January 2005 and 31 December 2006 with follow-up capturing all inpatient admissions to National Health Service (NHS) hospitals in England from birth until 31 March 2015. Children were not eligible if they had opted out, died before discharge from the birth admission or if there were data quality issues.⁸

The NNRD is a national resource holding quality-assured real-world clinical data captured during the course of care for all admissions to NHS neonatal units in England, Wales, Scotland and the Isle of Man.¹⁵ PICANet is an audit database recording demographic and clinical information on all patients admitted to paediatric intensive care units in the UK and Ireland.¹⁶

Study design

The main source of information for our economic analysis was the TIGAR dataset. However, although in the HES APC records the length of stay for admissions includes any time spent in critical care units, the information did not reliably indicate the level of care the child received on a day-to-day basis (online supplemental material 1.1). Therefore, we requested bespoke aggregate tables from the NNRD and PICANet to supplement the individual-level data within the TIGAR cohort. This information was used together with the total number of live births in England by gestational age in the same year, as reported by ONS, to simulate the following information for children in the TIGAR cohort by sex and gestational age at birth at an individual level²⁰:

- 1. Whether a child was admitted to a neonatal or paediatric critical care unit during the birth admission.
- 2. The number of days that were spent in a neonatal or paediatric critical care unit during the birth admission by level of care.

Any difference between the total birth admission days observed in the TIGAR cohort and the critical care days estimated from the NNRD and PICANet was considered non-critical care ward days. More details about the methods used to simulate neonatal and paediatric critical care days and calculate non-critical care ward days in birth admission can be found in online supplemental materials 1.2 and 1.3.

Costs

We estimated direct costs from a healthcare perspective in England. In the HES APC, each data record indicates a Finished Consultant Episode, which represents a continuous period of care under one clinical consultant. Costs were estimated at episode level. The 2018–2019 Casemix Grouper Software (HRG4+) was used to allocate each episode to a Healthcare Resource Group (HRG), primarily based on any procedures performed, diagnoses, hospital admission type, episode length of stay and patient characteristics.²¹ HRGs are standard groupings of clinically similar treatments, which use comparable levels of healthcare resource. The NHS 2017–2018 reference cost schedules were used to price the HRGs.²² More details on the application of the Grouper Software and matching the HRGs to reference costs can be found in online supplemental materials 2 and 3.

Neonatal and paediatric critical care costs during the birth admission were calculated on a per diem basis using the NHS 2017–2018 reference costs based on level of care. Birth admission ward costs were adjusted accordingly by extracting days spent in critical care units from the total hospital stay (online supplemental material 1.3).

Statistical analysis

We used descriptive statistics to estimate the average total length of stay and average costs for the birth admission and for subsequent hospital admissions by year of follow-up among children alive at the beginning and not censored by the end of each follow-up year. Comparisons were made between the following groups by gestational age at birth (weeks): <28, 28–29, 30–31, 32, 33, 34, 35, 36, 37, 38, 39, 40, 41, 42. Gestational ages were grouped <28, 28–29 and 30–31 following ONS policy about reporting of small numbers.

We used the Kaplan-Meier sample average estimator to calculate the total 1-year, 5-year and 8-year costs for each gestational age group.²³ The 95% CIs for total costs and p values for cost differences between gestational age at birth using 40 weeks as the reference were obtained using non-parametric bootstrapping with 1000 replications. No adjustment based on baseline characteristics was conducted.

In the baseline analysis, we calculated the birth admission non-critical care ward days and costs according to aggregated data from the NNRD and PICANet. We conducted a sensitivity analysis to investigate the results when allocating all the estimated non-critical care ward days to neonatal critical care for children born at ≤ 33 weeks to reflect an scenario where these babies spent most of their stay in a neonatal unit as previously indicated.²⁴

All analyses were conducted using Stata V.14 (College Station, Texas, USA). Costs are presented in 2018 UK pounds (£).

RESULTS

A total of 1018136 children were included in this study, with a total of 9372105 person years of follow-up and 1315338 admissions that occurred between 1 January 2005 and 31 March 2015. The baseline characteristics of the TIGAR cohort are presented in table 1. Among the 1018136 children included in this study, 56053 (5.5%) were born at <37 weeks, and 99717

Table 1	Baseline characteristics of live born singletons (total
N=.1018	136)

	n	%
Mother's age at childbirth (years)	11	/0
	44 486	4.4
20-24	181 633	17.8
25-29	253 055	24.9
30-34	293 741	24.9
35-39	193622	19.0
40+	51 599	5.1
Parity	100 64 6	47.0
Nulliparous	480616	47.2
Parous	496 203	48.7
Missing	41 317	4.1
Mother's marital status		
Married	581 160	57.1
Partner	347 366	34.1
Single	89610	8.8
Mother's country of birth		
Non-UK	225695	22.2
UK	791 012	77.7
Missing	1429	0.1
Index of Multiple Deprivation*		
Least deprived Q1	276838	27.2
Q2	216006	21.2
Q3	180300	17.7
Q4	161 793	15.9
Most deprived Q5	157195	15.4
Missing	26004	2.6
Sex		
Male	521 169	51.2
Female	496 967	48.8
Ethnicity (child)		
White British	677236	66.5
White other	59683	5.9
Bangladeshi	14546	1.4
Indian	27783	2.7
Pakistani	41 739	4.1
Black African	34571	3.4
Black Caribbean	12410	1.2
Other	91 570	9.0
Missing	58 598	5.8
Gestational age (weeks)		
<28	1730	0.2
28–29	2089	0.2
30-31	3227	0.3
32	2656	0.3
33	4050	0.4
34	7292	0.4
35	11 663	1.1
36	23 346	2.3
37	54001	5.3
38	137 926	13.5
39	231 376	22.7
40	288 065	28.3
41	208757	20.5
42	41 958	4.1
Small for gestational age at birth		

Table 1 Continued

	n	%
No	918419	90.2
Yes	99717	9.8
Mode of birth		
Vaginal	751 653	73.8
Caesarean section	222 615	21.9
Missing	43 868	4.3
Labour induction		
No	626178	61.5
Yes	154851	15.2
Missing	237107	23.3
Season of birth		
Jan–Mar	236944	23.3
Apr–Jun	254016	24.9
Jul–Sep	270282	26.5
Oct–Dec	256894	25.2
*The Index of Multiple Deprivation (IMD)	is the official measure of	rolativo

*The Index of Multiple Deprivation (IMD) is the official measure of relative deprivation for small areas (or neighbourhoods) in England.

(9.8%) were small for gestational age at birth (birth weight below the 10th centile).

The proportions of children admitted to neonatal or paediatric critical care during the birth admission are presented in table 2. The admission rates decreased with increasing gestational age at birth, with the lowest rate observed at 39–40 weeks (table 2). Children born extremely preterm (<28 weeks) spent. on average, 50.2 (SD 18.8) days in neonatal critical care, which generated an average cost of £53144 (SD 15504); in comparison, 8% of children born at 40 weeks were admitted to neonatal critical care with an average stay of 4.0 (SD 2.5) days, which generated an average cost of £2369 (SD £1783). Further, 5.5% of children born extremely preterm and 0.1% of children born at 40 weeks were admitted to paediatric critical care during their birth admission, which generated an average cost of £14967 (SD £9526) and £21957 (SD £14523), respectively. The average hospital days and costs incurred as a result of admission to noncritical care wards also decreased with increasing gestational age at birth (table 3), with the longest stay and highest cost observed for the extremely preterm group (48.8 days, £13142) and the shortest stay and lowest cost observed for those born at 40 gestational weeks (1.3 days, £244). The total costs associated with birth admissions including critical and non-critical care is presented in table 4.

Full details on estimated lengths of stay for admissions subsequent to the birth admission, by year of follow-up, can be found in the online supplemental materials 4. Associated average hospital costs for these subsequent admissions are presented in table 4. Similar to birth admission costs, subsequent hospital admission costs decreased with increasing gestational age at birth, with the largest differences between gestational age groups observed during the first year after birth. Differences in hospital costs between gestational age groups diminished with increasing age, particularly during the first 2 years after birth, while children born extremely preterm (<28 weeks) and very preterm (28–31 weeks) still had higher hospital admission costs (£699 (95% CI £419 to £919) for <28 weeks; £434 (95% CI £305 to £563) for 28–31 weeks) during the eighth year of life compared with children born at 40 weeks (£109, 95% CI £104 to £114).

The mean cumulative hospital admission cost over 8 years after birth among children born extremely preterm was estimated at

Continued

	% to NCC	Days in NCC, mean (SD)	Costs in NCC, mean (SD)	% to PCC	Days in PCC, mean (SD)	Costs in PCC, mean (SD)
<28	100%	50.2 (18.8)	£53 144 (15 504)	5.5%	8.8 (5.9)	£14967 (9526)
28–29	100%	39.3 (13.8)	£32 717 (10185)	1.7%	9.6 (4.0)	£18265 (9166)
30–31	100%	30.2 (9.3)	£21 957 (6753)	1.7%	9.0 (4.0)	£17596 (8361)
32	100%	22.0 (6.2)	£15350 (4772)	0.4%	16.1 (6.5)	£28530 (10386)
33	100%	16.3 (4.4)	£11 065 (3516)	0.4%	12.5 (7.8)	£24168 (15387)
34	90%	11.9 (4.4)	£7883 (3277)	0.4%	13.3 (7.0)	£24390 (13400)
35	63%	7.7 (3.9)	£5018 (2981)	0.4%	13.5 (6.6)	£24087 (13943)
36	40%	5.7 (3.4)	£3757 (2658)	0.4%	11.2 (6.8)	£20069 (13466)
37	20%	5.0 (2.9)	£3343 (2419)	0.2%	9.4 (5.0)	£19003 (11033)
38	11%	4.3 (2.8)	£2821 (2302)	0.2%	9.5 (4.9)	£19225 (11420)
39	8%	4.0 (2.5)	£2417 (1798)	0.1%	10.6 (5.9)	£23720 (15050)
40	8%	4.0 (2.5)	£2369 (1783)	0.1%	9.8 (5.9)	£21 957 (14 523)
41	10%	4.0 (2.4)	£2411 (1721)	0.1%	5.3 (3.7)	£13 080 (12 786)
42	12%	4.1 (2.6)	£2383 (1825)	0.1%	5.5 (3.2)	£10936 (9374)

 Table 2
 Proportion of children admitted to neonatal critical care (NCC) and paediatric critical care (PCC) during the birth admission and, among those admitted, length of stay and cost, by gestational age at birth

£80559 (95% CI £79238 to £82019) per child, with most of the cost incurred during the first year after birth (£71997 (95% CI £70866 to £73097)) (table 5). Children born at 40 weeks' gestational age incurred the lowest 1-year, 5-year and 8-year cumulative hospital admission costs compared with other gestational age groups, with all cost differences being statistically significant (table 5). Even children born at 39 weeks had a higher 8-year cumulative hospital admission cost (£2085 (95% CI £2061 to £2107)) compared with those born at 40 weeks (£1894 (95%) CI £1874 to £1912)). In the sensitivity analysis, after allocating all the non-critical care ward days to neonatal critical care for children born at \leq 33 weeks, estimated total hospital admission costs were higher at these gestational ages, especially for the extremely preterm group where the 8-year total hospital admission cost increased to £119044 (95% CI £117350 to £120909) per child.

DISCUSSION

In this study, we have investigated hospital admission costs from birth to mid-childhood in England across the full range of gestational age at birth. We found that gestational age at

Table 3Non-critical ward days and costs during the birthadmission, by gestational age at birth								
	n	Mean days in non-critical care ward	Mean costs in non-critical care ward	95% CI				
<28	1730	48.8	£13142	£12512 to £13771				
28–29	2089	23.4	£5567	£5310 to £5823				
30–31	3227	9.7	£2394	£2291 to £2498				
32	2656	5.4	£1377	£1311 to £1443				
33	4050	3.7	£1016	£963 to £1068				
34	7292	1.8	£532	£509 to £556				
35	11 663	2.4	£809	£775 to £843				
36	23346	2.2	£753	£725 to £781				
37	54001	1.8	£463	£446 to £480				
38	137 926	1.8	£373	£365 to £382				
39	231 376	1.6	£285	£279 to £290				
40	288065	1.3	£244	£240 to £249				
41	208757	1.4	£262	£258 to £267				
42	41 958	1.5	£342	£329 to £356				

birth is associated not only with birth admission hospital costs but also subsequent hospital admission costs up to age 8. The most common cause of subsequent admissions was infection.⁸ Children born extremely preterm were estimated to have high hospital admission costs throughout the first 8 years of life, with the majority of the costs incurred during the first year after birth.

Our cost estimates are in line with other studies of the costs of preterm birth in England. Khan *et al* reported similar birth admission hospital costs for 32-33 week moderately preterm born (£13 168) and 34-36 week late preterm born (£5463) children (2017–18 prices), based on a cohort in the East Midlands region of England.¹⁴ Mangham *et al* estimated the costs to the public sector over the first 18 years after birth using a decision-analytic model and reported neonatal care costs of £109 860 (2017–2018 price) for extremely preterm children.¹¹ This is higher than our base case estimates (£53 144) when 2017 NNRD data are applied, but similar to the costs generated by our sensitivity analysis (£110482) that assumed that the entire birth admission of children born at ≤ 33 weeks in TIGAR was spent in critical care.

This is, to our knowledge, the first study that uses a large national cohort to examine the association between gestational age at birth and hospital admission costs across the full spectrum of gestational age. The results of this study provide comprehensive estimates of hospital admission costs from birth to midchildhood by gestational age at birth, which can act as a useful resource for clinical and budgetary service planning, and as data inputs for economic evaluations of preventive and treatment strategies for preterm birth.

The main strength of this study is the large sample size available for analysis, which provided a sufficient sample to estimate costs across the full spectrum of gestational age. The national coverage of the TIGAR cohort contrasts with the regional-based or clinic-based populations that provided a focus for previous studies.^{10 14} The linkage to routinely collected HES data accompanied by standardised costing approaches through HRG Groupers ensured an accurate estimation of hospital costs, as it based the estimation on more granular patient-level hospital activities and nationally recommended cost algorithms.

The main limitation of the study is that we were not able to estimate more granular neonatal and paediatric critical care costs for the TIGAR cohort as such estimation is not supported

		Birth admission	Year 1	Year 2	Year 3	Year 4	Year 5	Year 6	Year 7	Year 8
<28	n	1730	1730	1710	1705	1701	1697	1695	1695	1691
	Mean	£67168	£4885	£2726	£1539	£1347	£933	£848	£656	£669
	SD	£33071	£10937	£11342	£6446	£8032	£4895	£7961	£4333	£5248
8–29	n	2089	2089	2071	2069	2067	2067	2066	2065	2064
	Mean	£38616	£2893	£1410	£716	£703	£567	£587	£564	£522
	SD	£18893	£7610	£6564	£3033	£4419	£4252	£4289	£5511	£4719
0-31	n	3227	3227	3218	3217	3214	3213	3212	3211	3211
	Mean	£24640	£2029	£754	£636	£421	£527	£434	£330	£377
	SD	£12231	£5698	£3309	£5178	£2225	£4082	£4119	£2142	£4805
2	n	2656	2656	2646	2645	2644	2642	2642	2641	2639
	Mean	£16803	£1813	£842	£592	£539	£442	£370	£344	£240
	SD	£8056	£6004	£4578	£3955	£4104	£2837	£2899	£2094	£1267
3	n	4050	4050	4026	4021	4018	4017	4016	4015	4015
	Mean	£12212	£1387	£667	£399	£311	£325	£256	£232	£194
	SD	£7316	£5931	£4234	£2819	£1695	£2733	£1571	£1458	£1218
4	n	7292	7292	7255	7247	7244	7242	7240	7240	7240
	Mean	£7732	£1399	£655	£396	£338	£251	£278	£262	£209
	SD	£6570	£7269	£5340	£3547	£3224	£1829	£2982	£2316	£1755
5	n	11663	11 663	11614	11 601	11 593	11 590	11 588	11 588	11 583
	Mean	£4052	£1086	£503	£311	£275	£259	£243	£226	£218
	SD	£6990	£5883	£4195	£2625	£2636	£2759	£2801	£2622	£2701
6	n	23346	23 346	23273	23255	23 249	23246	23239	23237	23 235
	Mean	£2325	£1030	£437	£314	£264	£279	£242	£205	£181
	SD	£6235	£6061	£3141	£3468	£2378	£3700	£2913	£2901	£2454
7	n	54001	54001	53 889	53 865	53 845	53 836	53 824	53818	53812
	Mean	£1166	£744	£356	£247	£233	£201	£196	£167	£159
	SD	£4709	£4012	£2953	£2296	£2516	£1921	£2482	£1677	£1887
8	n	137926	137926	137711	137654	137616	137 596	137 575	137 563	13755
	Mean	£725	£570	£291	£208	£191	£178	£164	£147	£143
	SD	£3839	£3438	£2347	£2079	£1990	£1765	£1642	£1608	£1952
9	n	231 376	231 376	231150	231 072	231 028	230 993	230970	230 950	23093
	Mean	£502	£453	£239	£175	£160	£160	£148	£127	£121
	SD	£3058	£2962	£2157	£1800	£1569	£1703	£1743	£1484	£1612
0	n	288 065	288 065	287 821	287 748	287691	287 663	287 644	287627	28760
	Mean	£455	£394	£227	£164	£149	£142	£133	£122	£109
	SD	£2860	£2690	£2071	£1899	£1496	£1412	£1401	£1467	£1352
1	n	208757	208757	208576	208 528	208 501	208 480	208 460	208 447	20843
	Mean	£519	£358	£219	£158	£151	£154	£135	£120	£115
	SD	£2585	£2419	£1898	£1614	£1494	£1926	£1582	£1426	£1519
2	n	41 958	41 958	41 917	41 902	41 896	41 890	41 888	41 881	41 880
	Mean	£641	£391	£224	£159	£168	£152	£131	£118	£116
	SD	£2849	£2935	£1702	£1360	£2056	£1582	£1302	£1523	£1637

The TIGAR cohort excludes babies who died within their birth admission. As a result, for subsequent admissions all babies were alive at the beginning of the first year.

by the HES APC data; notably, critical care information was collected with insufficient quality in HES APC before 2008. A separate HES dataset covers adult critical care from 2008/2009, whereas data relating to neonatal or paediatric intensive care are collected through systems external to NHS Digital, which collects HES data.¹⁷ To address this issue, we conducted an evidence synthesis exercise and simulated cost estimates for neonatal or paediatric intensive care using aggregated data from the NNRD and PICANet. This allowed us to account for critical care costs in the analysis. In the case of PICANet, aggregate data were provided only by gestational week bands instead of a specific week. Therefore, we had to assume that the information provided for a specific band was the same across all

weeks. This may have introduced some inaccuracies in our estimation of length of stay and associated costs for PCC. In addition, we were not able to obtain neonatal and paediatric critical care data over the same time coverage of the TIGAR cohort as such information does not date back to 2005–2006. This may have contributed to some of the differences between the total birth admission days observed in TIGAR and the neonatal and paediatric critical care days simulated using NNRD and PICANet data. Nevertheless, our sensitivity analysis that estimated costs using alternative assumptions about the ward stay during the birth admission of infants born either very or extremely preterm provides an upper bound for cost estimates for the birth admission.

 Table 5
 Total cumulative 1-year, 5-year and 8-year hospital cost (£) by gestational age at birth, estimated with the Kaplan-Meier sample-average estimator

	1-year tota	ıl		5-year total			8-year total		
	Mean	95% CI	P value	Mean	95% CI	P value	Mean	95% CI	P value
<28	71 997	70866 to 73097	<0.0001	78432	77 164 to 79 818	<0.0001	80559	79238 to 82019	<0.0001
28–29	41 484	40843, 42135	< 0.0001	44 846	44 097 to 45 584	< 0.0001	46 499	45 705, 47 291	< 0.0001
30–31	26663	26310 to 26989	< 0.0001	28 991	28551 to 29440	< 0.0001	30126	29612 to 30605	< 0.0001
32	18609	18313 to 18899	< 0.0001	21 013	20613 to 21447	<0.0001	21 961	21 550 to 22 409	< 0.0001
33	13 591	13 353 to 13 821	< 0.0001	15280	14979 to 15570	<0.0001	15955	15637 to16257	< 0.0001
34	9124	8918 to 9310	<0.0001	10755	10495 to 11007	<0.0001	11 498	11 214, 11 769	< 0.0001
35	5133	5001 to 5257	< 0.0001	6474	6303 to 6638	< 0.0001	7156	6967 to 7334	< 0.0001
36	3351	3262 to 3438	< 0.0001	4641	4523 to 4758	< 0.0001	5266	5131 to 5399	< 0.0001
37	1909	1866 to 1951	< 0.0001	2943	2883 to 3002	< 0.0001	3464	3402 to 3534	< 0.0001
38	1294	1273 to 1315	< 0.0001	2161	2128 to 2191	< 0.0001	2612	2576 to 2646	< 0.0001
39	955	940 to 970	< 0.0001	1689	1669 to 1710	< 0.0001	2085	2061 to 2107	< 0.0001
40	848	837 to 861	Reference	1530	1512 to 1547	Reference	1894	1874 to 1912	Reference
41	877	865 to 889	< 0.0001	1558	1540 to 1577	0.021	1928	1905 to 1950	0.013
42	1032	997 to 1066	< 0.0001	1732	1685 to 1777	< 0.0001	2096	2045 to 2148	< 0.0001
Sensitivity analysis*									
<28	110 482	108 946 to 112 149	< 0.0001	116917	115267 to 118675	< 0.0001	119044	117 350 to 120 909	< 0.0001
28–29	55365	54629 to 56171	< 0.0001	58727	57893 to 59638	< 0.0001	60380	59443 to 61404	< 0.0001
30–31	31 332	30939 to 31712	< 0.0001	33 660	33 191 to 34 136	< 0.0001	34794	34273 to 35336	< 0.0001
32	20982	20656 to 21303	<0.0001	23 386	22 918 to 23 834	<0.0001	24334	23838 to 24804	< 0.0001
33	15079	14845 to 15320	<0.0001	16 768	16451 to 17059	<0.0001	17444	17125 to 17747	< 0.0001

It is also worth highlighting that the TIGAR cohort excluded children who died before discharge from their birth admission, which means that we have excluded a small group of extremely ill babies from our study. This suggests that our cost estimates should be viewed as conservative, particularly for the purposes of planning of neonatal services.

In conclusion, this study provides estimates of the association between gestational age at birth and hospital admission costs from birth to mid-childhood and disaggregates those estimates by gestational category and chronological year. The study results should act as a useful resource for future economic evaluations that focus on preventive and treatment strategies for preterm birth and inform resource allocation decisions.

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Acknowledgements We are grateful to the following organisations and individuals who were involved in this study: Medical Research Council (MRC) for funding this project (MR/M01228X/1); Office for National Statistics (ONS) for providing access to data and hosting it within the safe research setting (SRS); NHS Digital for linking and providing access to the data; National Neonatal Research Database (NNRD) for providing aggregated data on neonatal critical care admissions; Paediatric Intensive Care Audit Network (PICANet) for providing aggregated data on paediatric critical care admissions; the TIGAR patient and public involvement (PPI) group and National Childbirth Trust (NCT) representatives for their input into the study objectives and design; and the TIGAR advisory group for their input into the study design and dissemination. This work contains statistical data in this work does not imply endorsement of the ONS of the interpretation or analysis of the statistical data. This work uses research datasets, which might not exactly reproduce National Statistics aggregates.

Contributors CRediT (Contributor Roles Taxonomy) author statement: Conceptualisation: OR-A, SP, MAQ, JJK. Methodology: XH, OR-A, SP. Formal analysis: XH with input from OR-A, CC and VC. Data curation: MAQ, AM. Writing—Original draft: XH with input from OR-A and SP. Writing—Review and editing: XH, SP, VC, CC, JJK, MAQ, EB, SJ, AM, OR-A. Visualisation: XH, OR-A. Supervision: OR-A. Project administration: OR-A, MAQ. Funding acquisition: MAQ, OR-A, SP, CC, JJK, EB, SJ, AM. OR-A acts as the guarantor. **Funding** The TIGAR study was funded by a research grant from the MRC (MR/ M01228X/1). The funder had no role in the study design or in the collection, analysis, interpretation of data, writing of the report or decision to submit the article for publication. XH is an AUS NHMRC Emerging Leadership Fellow (grant number 2009253). SP receives support as a UK National Institute for Health Research (NIHR) Senior Investigator (NF-SI-0616–10103) and from the UK NIHR Applied Research Collaboration Oxford and Thames Valley.

Competing interests None declared.

Patient consent for publication Not applicable.

Ethics approval Ethical approval for this study was granted by the Health Research Authority research ethics committee (South West–Frenchay; REC reference 15/SW/0294) and was provided by the Health Research Authority following Advice from the Confidentiality Advisory Group (CAG reference 15/CAG/0196). Participants gave informed consent to participate in the study before taking part.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Data may be obtained from a third party and are not publicly available. The authors do not have permission to supply data or identifiable information to third parties, including other researchers, but the team at City, University of London has permission under regulation 5 of the Health Service (control of patient information) Regulations 2002 to analyse patient identifiable data for England and Wales without consent and create a research database that could be accessed by other researchers using the SRS at the ONS. The TIGAR team has permission under regulations 5 of the Health Service (control of patient information) Regulations 2002 to analyse these data. Anyone wishing to access the linked datasets for research purposes should apply via the CAG to the Health Research Authority to access patient identifiable data without consent and then to the ONS and NHS Digital. In the first instance, enquiries about access to the data should be addressed to Alison Macfarlane.

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REFERENCES

- 1 Goldenberg RL, Culhane JF, Iams JD, *et al*. Epidemiology and causes of preterm birth. *Lancet* 2008;371:75–84.
- 2 Blencowe H, Cousens S, Oestergaard MZ, et al. National, regional, and worldwide estimates of preterm birth rates in the year 2010 with time trends since 1990 for selected countries: a systematic analysis and implications. Lancet 2012;379:2162–72.
- 3 Chawanpaiboon S, Vogel JP, Moller A-B, et al. Global, regional, and national estimates of levels of preterm birth in 2014: a systematic review and modelling analysis. *Lancet Glob Health* 2019;7:e37–46.
- 4 Stoll BJ, Hansen NI, Bell EF, *et al.* Trends in care practices, morbidity, and mortality of extremely preterm neonates, 1993-2012. *JAMA* 2015;314:1039–51.
- 5 Santhakumaran S, Statnikov Y, Gray D, et al. Survival of very preterm infants admitted to neonatal care in England 2008-2014: time trends and regional variation. Arch Dis Child Fetal Neonatal Ed 2018;103:F208–15.
- 6 Saigal S, Doyle LW. An overview of mortality and sequelae of preterm birth from infancy to adulthood. *Lancet* 2008;371:261–9.
- 7 Moster D, Lie RT, Markestad T. Long-Term medical and social consequences of preterm birth. N Engl J Med 2008;359:262–73.
- 8 Coathup V, Boyle E, Carson C, et al. Gestational age and hospital admissions during childhood: population based, record linkage study in England (TIGAR study). BMJ 2020;371:m4075.
- 9 Petrou S, Yiu HH, Kwon J. Economic consequences of preterm birth: a systematic review of the recent literature (2009-2017). Arch Dis Child 2019;104:456–65.

10 Petrou S. The economic consequences of preterm birth during the first 10 years of life. BJOG 2005;112 Suppl 1:10–5.

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- 11 Mangham LJ, Petrou S, Doyle LW, et al. The cost of preterm birth throughout childhood in England and Wales. *Pediatrics* 2009;123:e312–27.
- 12 Johnston KM, Gooch K, Korol E, *et al*. The economic burden of prematurity in Canada. *BMC Pediatr* 2014;14:93.
- 13 Petrou S, Abangma G, Johnson S, *et al*. Costs and health utilities associated with extremely preterm birth: evidence from the EPICure study. *Value Health* 2009;12:1124–34.
- 14 Khan KA, Petrou S, Dritsaki M, *et al*. Economic costs associated with moderate and late preterm birth: a prospective population-based study. *BJOG* 2015;122:1495–505.
- 15 Gale C, Morris I, Neonatal Data Analysis Unit (NDAU) Steering Board. The UK national neonatal research database: using neonatal data for research, quality improvement and more. Arch Dis Child Educ Pract Ed 2016;101:216–8.
- 16 Paediatric intensive care audit network annual report (2020) (published 02/2021): universities of leeds and leicester. online. 2020.
- 17 Herbert A, Wijlaars L, Zylbersztejn A, et al. Data resource profile: Hospital episode statistics admitted patient care (Hes APC). Int J Epidemiol 2017;46:1093–1093i.
- 18 Macfarlane A, Dattani N, Gibson R, et al. Births and their outcomes by time, day and year: a retrospective birth cohort data linkage study. *Health Serv Deliv Res* 2019;7:1–268.
- 19 Coathup V, Macfarlane A, Quigley M. Linkage of maternity hospital episode statistics birth records to birth registration and notification records for births in England 2005-2006: quality assurance of linkage. *BMJ Open* 2020;10:e037885.
- 20 Office for national statistics. Live births by gestational age, where mother's usual residence was england, 2016 and 2017. ref: 009545. 2019.
- 21 National Casemix Office. Grouper user manual: HRG4+ 2018/19 reference costs grouper. 2019.
- 22 NHS. National schedule of NHS costs 2017/18. 2018.
- 23 Lin DY, Feuer EJ, Etzioni R, *et al*. Estimating medical costs from incomplete follow-up data. *Biometrics* 1997;53:419–34.
- 24 Seaton SE, Barker L, Draper ES, *et al*. Estimating neonatal length of stay for babies born very preterm. *Arch Dis Child Fetal Neonatal Ed* 2019;104:F182–6.